#### Fundamentals of Artificial Intelligence Chapter 13: **Quantifying Uncertainty**

#### Roberto Sebastiani

DISI, Università di Trento, Italy - roberto.sebastiani@unitn.it http://disi.unitn.it/rseba/DIDATTICA/fai\_2020/

Teaching assistant: Mauro Dragoni - dragoni@fbk.eu http://www.maurodragoni.com/teaching/fai/

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### Outline

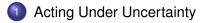


- Acting Under Uncertainty
- Basics on Probability 2
- Probabilistic Inference via Enumeration
- Independence and Conditional Independence
- Applying Bayes' Rule 5



6 An Example: The Wumpus World Revisited

#### Outline



- 2 Basics on Probability
- 3 Probabilistic Inference via Enumeration
- Independence and Conditional Independence
- 5 Applying Bayes' Rule
- 6 An Example: The Wumpus World Revisited

Agents often make decisions based on incomplete information

- partial observability
- nondeterministic actions
- Partial solution (see previous chapters): maintain belief states
   represent the set of all possible world states the agent might be in
   generating a contingency plan handling every possible eventuality
- Several drawbacks:
  - must consider every possible explanation for the observation (even very-unlikely ones) ⇒ impossibly complex belief-states
  - contingent plans handling every eventuality grow arbitrarily large
     sometimes there is no plan that is guaranteed to achieve the goal
- Agent's knowledge cannot guarantee a successful outcome ...
  - .. but can provide some degree of belief (likelihood) on it
- A rational decision depends on both the relative importance of (sub)goals and the likelihood that they will be achieved
- Probability theory offers a clean way to quantify likelihood

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#### Automated taxi to Airport

- Goal: deliver a passenger to the airport on time
- Action A<sub>t</sub>: leave for airport t minutes before flight
   How can we be sure that A<sub>90</sub> will succeed?
- Too many sources of uncertainty:
  - partial observability (ex: road state, other drivers' plans, etc.)
  - uncertainty in action outcome (ex: flat tire, etc.)
  - noisy sensors (ex: unreliable traffic reports)
  - complexity of modelling and predicting traffic
- With purely-logical approach it is difficult to anticipate everything that can go wrong
  - risks falsehood: "A25 will get me there on time" or
  - leads to conclusions that are too weak for decision making:
     "A<sub>25</sub> will get me there on time if there's no accident on the bridge and it doesn't rain and my tires remain intact , and..."

• Over-cautious choices are not rational solutions either

• ex: A<sub>1440</sub> causes staying overnight at the airport

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#### A medical diagnosis

- Given the symptoms (toothache) infer the cause (cavity)
- How to encode this relation in logic?
  - diagnostic rules:
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    - (too many possible causes, some very unlikely)
  - causal rules:
    - Cavity  $\rightarrow$  Toothache (wrong)
    - $(Cavity \land ...) \rightarrow Toothache$  (many possible (con)causes
- Problems in specifying the correct logical rules:
  - Complexity: too many possible antecedents or consequents
  - Theoretical ignorance: no complete theory for the domain
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- Probability allows to summarize the uncertainty on effects of
  - laziness: failure to enumerate exceptions, qualifications, etc.
  - ignorance: lack of relevant facts, initial conditions, etc.
- Probability can be derived from
  - statistical data (ex: 80% of toothache patients so far had cavities)
  - some knowledge (ex: 80% of toothache patients has cavities)
  - their combination thereof
- Probability statements are made with respect to a state of knowledge (aka evidence), not with respect to the real world
  - e.g., "The probability that the patient has a cavity, given that she has a toothache, is 0.8":
    - *P*(*HasCavity*(*patient*) | hasToothAche(patient)) = 0.8
- Probabilities of propositions change with new evidence:
  - "The probability that the patient has a cavity, given that she has a toothache and a history of gum disease, is 0.4": P(HasCavity(patient)

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• Ex: Suppose I believe:

- Depends on tradeoffs among preferences:
   missing flight vs. costs (airport cuisine, sleep overnight in airport)
- When there are conflicting goals the agent may express preferences among them by means of a utility function.
- Utilities are combined with probabilities in the general theory of rational decisions, aka decision theory: Decision theory = Probability theory + Utility theory
- Maximum Expected Utility (MEU): an agent is rational if and only if it chooses the action that yields the maximum expected utility, averaged over all the possible outcomes of the action.

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#### Probabilities Basics: an Al-sh Introduction

#### • Probabilistic assertions: state how likely possible worlds are

#### Sample space Ω: the set of all possible worlds

- $\omega \in \Omega$  is a possible world (aka sample point or atomic event)
- ex: the dice roll (1,4)
- the possible worlds are mutually exclusive and exhaustive
- ex: the 36 possible outcomes of rolling two dice: (1,1), (1,2), ...
- A probability model (aka probability space) is a sample space with an assignment P(ω) for every ω ∈ Ω s.t.
  - $0 \leq P(\omega) \leq 1$ , for every  $\omega \in \Omega$
  - $\Sigma_{\omega \in \Omega} P(\omega) = 1$
- Ex: 1-die roll: P(1) = P(2) = P(3) = P(4) = P(5) = P(6) = 1/6
- An Event A is any subset of  $\Omega$ , s.t.  $P(A) = \sum_{\omega \in A} P(\omega)$ 
  - events can be described by propositions in some formal language
  - ex: P(Total = 11) = P(5,6) + P(6,5) = 1/36 + 1/36 = 1/18
  - ex: P(doubles) = P(1,1) + P(2,2) + ... + P(6,6) = 6/36 = 1/6

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#### **Random Variables**

 Factored representation of possible worlds: sets of (*variable*, *value*) pairs

- Variables in probability theory: Random variables
  - domain: the set of possible values a variable can take on ex: Die: {1,2,3,4,5,6}, Weather: {*sunny*, *rain*, *cloudy*, *snow*}, Odd: {*true*, *false*},
  - a r.v. can be seen as a function from sample points to the domain:
     ex: Die(ω), Weather(ω),... ("(ω)" typically omitted)

• Probability Distribution gives the probabilities of all the possible values of a random variable *X*:  $P(X = x_i) \stackrel{\text{def}}{=} \sum_{\omega \in X(\omega)} P(\omega)$ 

• ex: P(Odd = true) = P(1) + P(3) + P(5) = 1/6 + 1/6 + 1/6 = 1/2

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- We think a proposition *a* as the event *A* (set of sample points) where the proposition is true
  - Odd is a propositional random variable of range {true, false}
  - notation:  $a \iff "A = true"$
- Given Boolean random variables A and B:
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- $\Rightarrow\,$  with Boolean random variables, sample points are PL models
- Proposition: disjunction of the sample points in which it is true
  - ex:  $(a \lor b) \equiv (\neg a \land b) \lor (a \land \neg b) \lor (a \land b)$
  - $\Rightarrow P(a \lor b) = P(\neg a \land b) + P(a \land \neg b) + P(a \land b)$
- Some derived facts:
  - $P(\neg a) = 1 P(a)$
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# **Probability Distributions**

- Probability Distribution gives the probabilities of all the possible values of a random variable
  - ex: Weather: {*sunny*, *rain*, *cloudy*, *snow*}
  - $\implies$  **P**(Weather) = (0.6, 0.1, 0.29, 0.01)  $\iff$ 

    - $\begin{array}{l} P(Weather = sunny) &= 0.6 \\ P(Weather = rain) &= 0.1 \\ P(Weather = cloudy) &= 0.29 \end{array}$
    - P(Weather = snow) = 0.01
    - ormalized: their sum is 1
- Joint Probability Distribution for multiple variables
  - gives the probability of every sample point
  - ex:  $\mathbf{P}(Weather, Cavity) =$

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Weather =	sunny	rain	cloudy	snow
Cavity = true	0.144	0.02	0.016	0.02
Cavity = false	0.576	0.08	0.064	0.08

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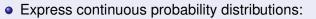
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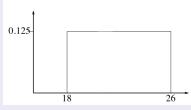


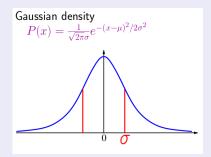
• density functions  $f(x) \in [0, 1]$  s.t  $\int_{-\infty}^{+\infty} f(x) dx = 1$ 

•  $P(x \in [a, b]) = \int_{a}^{b} f(x) dx$   $\implies P(x \in [val, val]) = 0, P(x \in [-\infty, +\infty]) = 1$ •  $ex: P(x \in [20, 22]) = \int_{20}^{22} 0.125 dx = 0.25$ 

• Density:  $P(x) = P(X = x) \stackrel{\text{def}}{=} \lim_{dx \mapsto 0} P(X \in [x, x + dx])/dx$ • ex:  $P(20.1) = \lim_{dx \mapsto 0} P(X \in [20.1, 20.1 + dx])/dx = 0.125$ • pote:  $P(x) \neq P(x \in [x, y]) = 0$ 

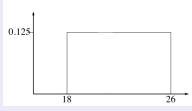
Uniform density between 18 and 26f(x) = U[18, 26](x)

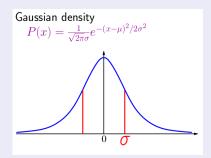




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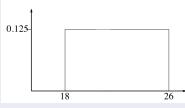
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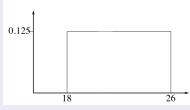
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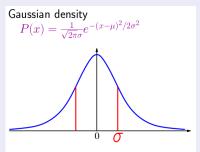
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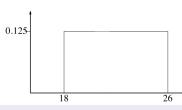
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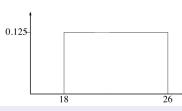
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 Unconditional or prior probabilities refer to degrees of belief in propositions in the absence of any other information (evidence)

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Less specific belief still valid after more evidence arrives
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New evidence may be irrelevant, allowing for simplification
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### Logic vs. Probability

$$\begin{array}{c|c} Logic & Probability \\ \hline a & P(a) = 1 \\ \neg a & P(a) = 0 \\ a \rightarrow b & P(b|a) = 1 \\ \hline (a, a \rightarrow b) & P(b|a) = 1 \\ \hline (a \rightarrow b, b \rightarrow c) \\ \hline a \rightarrow c & P(b|a) = 1, P(b|a) = 1 \\ \hline P(b|a) = 1, P(c|b) = 1 \\ \hline P(c|a) = 1 \end{array}$$

• Proof of P(b|a) = 1,  $P(c|b) = 1 \implies P(c|a) = 1$ 

- $P(b|a) = 1 \Longrightarrow P(\neg b, a) \stackrel{\text{def}}{=} P(\neg b|a)P(a) = 0$
- $P(c|b) = 1 \Longrightarrow P(\neg c, b) \stackrel{\text{def}}{=} P(\neg c|b)P(b) = 0$
- $P(\neg c, a) = P(\neg c, a, b) + P(\neg c, a, \neg b) \leq \underbrace{P(\neg c, b)}_{P(\neg c, b)} + \underbrace{P(a, \neg b)}_{P(\neg c, b)} = 0$

(Courtesy of Maria Simi, UniPI)

# Outline



### 2 Basics on Probability

### Probabilistic Inference via Enumeration

- Independence and Conditional Independence
- 5 Applying Bayes' Rule
- 6 An Example: The Wumpus World Revisited

## Probabilistic Inference via Enumeration

### **Basic Ideas**

Start with the joint distribution P(Toothache, Catch, Cavity)

• For any proposition  $\varphi$ , sum the atomic events where  $\varphi$  is true:  $P(\varphi) = \sum_{\omega : \omega \models \varphi} P(\omega)$ 

## Probabilistic Inference via Enumeration

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# Probabilistic Inference via Enumeration: Example

### Example: Generic Inference

- Start with the joint distribution P(Toothache, Catch, Cavity)
- For any proposition  $\varphi$ , sum the atomic events where  $\varphi$  is true:  $P(\varphi) = \sum_{\omega : \omega \models \varphi} P(\omega)$ :

 Ex: P(cavity ∨ toothache) = 0.108 + 0.012 + 0.072 + 0.008 + 0.016 + 0.064 = 0.28

	toothache		<i>¬ toothache</i>	
	catch	$\neg$ catch	catch	$\neg$ catch
cavity	.108	.012	.072	.008
$\neg$ cavity	.016	.064	.144	.576

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# Marginalization

• Start with the joint distribution **P**(*Toothache*, *Catch*, *Cavity*)

• Marginalization (aka summing out): sum up the probabilities for each possible value of the other variables:

 $\mathsf{P}(\mathsf{Y}) = \sum_{\mathsf{z}\in\mathsf{Z}}\mathsf{P}(\mathsf{Y},\mathsf{z})$ 

Ex:  $P(Toothache) = \sum_{z \in \{Catch, Cavity\}} P(Toothache, z)$ 

Conditioning: variant of marginalization, involving conditional probabilities instead of joint probabilities (using the product rule)
 P(Y) = ∑<sub>z∈Z</sub> P(Y|z)P(z)
 Ex: P(Toothache) = ∑ are as a P(Toothache|z)P(z)

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  - Ex:  $P(\text{Toothache}) = \sum_{z \in \{\text{Catch, Cavity}\}} P(\text{Toothache}, z)$
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Ex:  $\mathbf{P}(\text{Toothache}) = \sum_{\mathbf{z} \in \{\text{Catch}, \text{Cavity}\}} \mathbf{P}(\text{Toothache}|\mathbf{z}) P(\mathbf{z})$ 

# Marginalization: Example

- Start with the joint distribution P(Toothache, Catch, Cavity)
- Marginalization (aka summing out): sum up the probabilities for each possible value of the other variables:

 $\begin{aligned} \mathbf{P}(\mathbf{Y}) &= \sum_{\mathbf{z} \in \mathbf{Z}} \mathbf{P}(\mathbf{Y}, \mathbf{z}) \\ \text{Ex: } \mathbf{P}(\text{Toothache}) &= \sum_{\mathbf{z} \in \{\text{Catch}, \text{Cavity}\}} \mathbf{P}(\text{Toothache}, \mathbf{z}) \\ P(\text{toothache}) &= 0.108 + 0.012 + 0.016 + 0.064 = 0.2 \\ P(\neg \text{toothache}) &= 1 - P(\text{toothache}) = 1 - 0.2 = 0.8 \end{aligned}$ 

 $\implies$  **P**(*Toothache*) =  $\langle 0.2, 0.8 \rangle$ 

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# Conditional Probability via Enumeration: Example

- Start with the joint distribution P(Toothache, Catch, Cavity)
- Conditional Probability:

Ex:  $P(\neg cavity | toothache) = \frac{P(\neg cavity \land toothache)}{P(toothache)}$ =  $\frac{0.016+0.064}{0.108+0.012+0.016+0.064} = 0.4$ Ex:  $P(cavity | toothache) = \frac{P(cavity \land toothache)}{P(toothache)} = ... = 0.4$ 

	toothache		<i>¬ toothache</i>	
	catch	$\neg$ catch	catch	$\neg$ catch
cavity	.108	.012	.072	.008
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Ex: *P*(*cavity*|*toothache*) =

$$\frac{P(cavity \land toothache)}{P(toothache)} = \dots = 0.$$

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	catch	$\neg$ catch	catch	$\neg$ catch
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#### Normalization

• Let **X** be all the variables. Typically, we want P(Y|E = e):

- the conditional joint distribution of the query variables Y
- given specific values e for the evidence variables E
- let the hidden variables be  $\mathbf{H} \stackrel{\text{def}}{=} \mathbf{X} \setminus (\mathbf{Y} \cup \mathbf{E})$
- The summation of joint entries is done by summing out the hidden variables:

 $\mathsf{P}(Y|\mathsf{E}=\mathsf{e}) = lpha\mathsf{P}(Y,\mathsf{E}=\mathsf{e}) = lpha\Sigma_{\mathsf{h}\in\mathsf{H}}\mathsf{P}(Y,\mathsf{E}=\mathsf{e},\mathsf{H}=\mathsf{h})$ 

where  $\alpha \stackrel{\text{def}}{=} 1/\mathbf{P}(\mathbf{E} = \mathbf{e})$  (different  $\alpha$ 's for different values of  $\mathbf{e}$ )

 $\Rightarrow$  it is easy to compute lpha by normalization

 note: the terms in the summation are joint entries, because Y, E, H together exhaust the set of random variables X

• Complexity:  $O(2^n)$ , *n* number of propositions  $\implies$  impractical

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# Normalization: Example

- $\alpha \stackrel{\text{def}}{=} 1/P(toothache)$  can be viewed as a normalization constant
- Idea: compute whole distribution on query variable by:
  - fixing evidence variables and summing over hidden variables
  - normalize the final distribution, so that  $\sum ... = 1$

Ex:

 $\mathbf{P}(Cavity | toothache) = \alpha \mathbf{P}(Cavity \land toothache)$ 

 $= \alpha [\mathbf{P}(Cavity, toothache, catch) + \mathbf{P}(Cavity, toothache, \neg catch)]$ =  $\alpha [\langle 0.108, 0.016 \rangle + \langle 0.012, 0.064 \rangle]$ 

 $= \alpha \langle 0.12, 0.08 \rangle = (\textit{normalization}) = \langle 0.6, 0.4 \rangle [\alpha = 5]$ 

 $\mathbf{P}(Cavity | \neg toothache) = ... = \alpha \langle 0.08, 0.72 \rangle = \langle 0.1, 0.9 \rangle [\alpha = 1.25]$ 

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#### Ex:

- $P(Cavity | toothache) = \alpha P(Cavity \land toothache)$
- $= \alpha$ [**P**(*Cavity*, *toothache*, *catch*) + **P**(*Cavity*, *toothache*,  $\neg$ *catch*)]
- $= \alpha [\langle 0.108, 0.016 \rangle + \langle 0.012, 0.064 \rangle]$
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 $\mathbf{P}(\textit{Cavity} | \neg \textit{toothache}) = ... = \alpha \langle 0.08, 0.72 \rangle = \langle 0.1, 0.9 \rangle [\alpha = 1.25]$ 

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- $= \alpha [\langle 0.108, 0.016 \rangle + \langle 0.012, 0.064 \rangle]$
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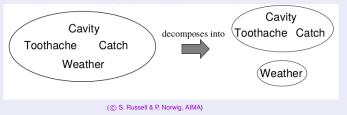
## Outline



- 2 Basics on Probability
- 3 Probabilistic Inference via Enumeration
- Independence and Conditional Independence
- 5 Applying Bayes' Rule
- 6 An Example: The Wumpus World Revisited

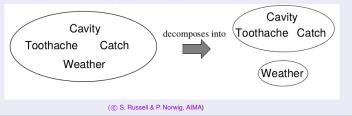
#### Independence

- Variables X and Y are independent iff P(X, Y) = P(X)P(Y) (or equivalently, iff P(X|Y) = P(X) or P(Y|X) = P(Y))
  - ex: P(Toothache, Catch, Cavity, Weather) = P(Toothache, Catch, Cavity)P(Weather)
  - $\implies$  e.g. P(toothache, catch, cavity, cloudy) = P(toothache, catch, cavity)P(cloudy)
    - typically based on domain knowledge
- May drastically reduce the number of entries and computation
   ex: 32-element table decomposed into one 8-element and one 4-element table
- Unfortunately, absolute independence is quite rare



#### Independence

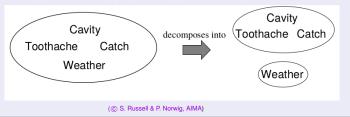
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    - typically based on domain knowledge
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  - ex: 32-element table decomposed into one 8-element and one 4-element table
- Unfortunately, absolute independence is quite rare



#### **Conditional Independence**

 Variables X and Y are conditionally independent given Z iff  $\mathbf{P}(X, Y|\mathbf{Z}) = \mathbf{P}(X|\mathbf{Z})\mathbf{P}(Y|\mathbf{Z})$ (or equivalently, iff P(X|Y,Z) = P(X|Z) or P(Y|X,Z) = P(Y|Z)) Consider P(Toothache, Cavity, Catch) if I have a cavity, the probability that the probe catches in it doesn't • the same independence holds if I haven't got a cavity:

#### **Conditional Independence**

- Variables X and Y are conditionally independent given Z iff
   P(X, Y|Z) = P(X|Z)P(Y|Z)
   (or equivalently, iff P(X|Y,Z) = P(X|Z) or P(Y|X,Z) = P(Y|Z))
- Consider P(Toothache, Cavity, Catch)
  - if I have a cavity, the probability that the probe catches in it doesn't depend on whether I have a toothache:
    - P(catch|toothache, cavity) = P(catch|cavity)
  - the same independence holds if I haven't got a cavity:  $P(catch|toothache, \neg cavity) = P(catch|\neg cavity)$
  - $\implies \mbox{Catch is conditionally independent of Toothache given Cavity:} \mbox{P}(Catch|Toothache, Cavity) = \mbox{P}(Catch|Cavity) \mbox{ or, equivalently:}$ 
    - P(Toothache|Catch, Cavity) = P(Toothache|Cavity), or
    - P(Toothache, Catch|Cavity) =
    - **P**(Toothache|Cavity)P(Catch|Cavity)

- In many cases, the use of conditional independence reduces the size of the representation of the joint distribution dramatically
  - even from exponential to linear!
- Ex: = P(Toothache|Catch, Cavity)P(Catch, Cavity) = P(Toothache|Catch, Cavity)P(Catch|Cavity)P(Cavity) • P(Toothache, Catch, Cavity) contains 7 independent entries P(Toothache|Cavity),P(Catch|Cavity) contain 2 independent • **P**(*Cavity*) contains 1 independent entry
- General Case: if one causes has n independent effects:  $P(Cause, Effect_1, ..., Effect_n) = P(Cause) \prod_i P(Effect_i | Cause)$

 $\implies$  reduces from  $2^{n+1} - 1$  to 2n + 1 independent entries

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- P(Toothache, Catch, Cavity)
- Ex: = P(Toothache|Catch, Cavity)P(Catch, Cavity)
  - $= \mathbf{P}(Toothache|Catch, Cavity)\mathbf{P}(Catch|Cavity)\mathbf{P}(Cavity)$
  - $= \mathbf{P}(Toothache|Cavity)\mathbf{P}(Catch|Cavity)\mathbf{P}(Cavity)$

Passes from 7 to 2+2+1=5 independent numbers

- P(Toothache, Catch, Cavity) contains 7 independent entries (the 8th can be obtained as 1 − ∑ ...)
- P(Toothache|Cavity),P(Catch|Cavity) contain 2 independent entries (2 × 2 matrix, each row sums to 1)
- **P**(*Cavity*) contains 1 independent entry

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 In many cases, the use of conditional independence reduces the size of the representation of the joint distribution dramatically

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- P(Toothache, Catch, Cavity)
- Ex:  $= \mathbf{P}(Toothache|Catch, Cavity)\mathbf{P}(Catch, Cavity)$ 
  - $= \mathbf{P}(Toothache|Catch, Cavity)\mathbf{P}(Catch|Cavity)\mathbf{P}(Cavity)$
  - $= \mathbf{P}(Toothache|Cavity)\mathbf{P}(Catch|Cavity)\mathbf{P}(Cavity)$

 $\implies$  Passes from 7 to 2+2+1=5 independent numbers

- P(*Toothache*, *Catch*, *Cavity*) contains 7 independent entries (the 8th can be obtained as 1 − ∑ ...)
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• General Case: if one causes has n independent effects:  $P(Cause, Effect_1, ..., Effect_n) = P(Cause) \prod_i P(Effect_i | Cause)$   $\rightarrow$  reduces from  $2^{n+1} = 1$  to  $2n \pm 1$  independent entries

 In many cases, the use of conditional independence reduces the size of the representation of the joint distribution dramatically

even from exponential to linear!

- **P**(Toothache, Catch, Cavity)
- Ex: = P(Toothache|Catch, Cavity)P(Catch, Cavity) = P(Toothache|Catch, Cavity)P(Catch|Cavity)P(Cavity)
  - $= \mathbf{P}(Toothache|Cavity)\mathbf{P}(Catch|Cavity)\mathbf{P}(Cavity)$

 $\implies$  Passes from 7 to 2+2+1=5 independent numbers

- P(Toothache, Catch, Cavity) contains 7 independent entries (the 8th can be obtained as  $1 - \sum ...$ )
- P(Toothache|Cavity), P(Catch|Cavity) contain 2 independent entries  $(2 \times 2 \text{ matrix}, \text{ each row sums to } 1)$
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- General Case: if one causes has n independent effects:  $P(Cause, Effect_1, ..., Effect_n) = P(Cause) \prod_i P(Effect_i | Cause)$

 $\implies$  reduces from  $2^{n+1} - 1$  to 2n + 1 independent entries

#### Exercise

Consider the joint probability distribution described in the table in previous section (slide 20 onwards): **P**(*Toothache*, *Catch*, *Cavity*)

- Consider the example in previous slide:
  - P(Toothache, Catch, Cavity)
  - $= \mathbf{P}(Toothache|Catch, Cavity)\mathbf{P}(Catch, Cavity)$
  - = **P**(*Toothache*|*Catch*, *Cavity*)**P**(*Catch*|*Cavity*)**P**(*Cavity*)
  - $= \mathbf{P}(Toothache|Cavity)\mathbf{P}(Catch|Cavity)\mathbf{P}(Cavity)$
- Compute separately the distributions
   P(Toothache|Catch, Cavity), P(Catch|Cavity), P(Cavity),
   P(Toothache|Cavity).

• Recompute **P**(*Toothache*, *Catch*, *Cavity*) in two ways:

- **P**(Toothache|Catch, Cavity)**P**(Catch|Cavity)**P**(Cavity)
- P(Toothache|Cavity)P(Catch|Cavity)P(Cavity)

and compare the result with P(Toothache, Catch, Cavity)

#### Exercise

Consider the joint probability distribution described in the table in previous section (slide 20 onwards): **P**(*Toothache*, *Catch*, *Cavity*)

- Consider the example in previous slide:
  - P(Toothache, Catch, Cavity)
  - $= \mathbf{P}(Toothache|Catch, Cavity)\mathbf{P}(Catch, Cavity)$
  - = P(Toothache|Catch, Cavity)P(Catch|Cavity)P(Cavity)
  - = **P**(*Toothache*|*Cavity*)**P**(*Catch*|*Cavity*)**P**(*Cavity*)
- Compute separately the distributions
   P(Toothache|Catch, Cavity), P(Catch|Cavity), P(Cavity),
   P(Toothache|Cavity).

• Recompute **P**(*Toothache*, *Catch*, *Cavity*) in two ways:

- **P**(Toothache|Catch, Cavity)**P**(Catch|Cavity)**P**(Cavity)
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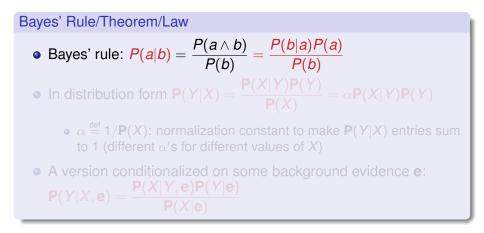
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### Outline



- 2 Basics on Probability
- Probabilistic Inference via Enumeration
- Independence and Conditional Independence
- Applying Bayes' Rule
  - An Example: The Wumpus World Revisited

#### Bayes' Rule



#### Bayes' Rule

#### Bayes' Rule/Theorem/Law

Bayes' rule: P(a|b) = P(a ∧ b)/P(b) = P(b|a)P(a)/P(b)
In distribution form P(Y|X) = P(X|Y)P(Y)/P(Y)/P(X) = αP(X|Y)P(Y)

α <sup>def</sup> = 1/P(X): normalization constant to make P(Y|X) entries sum to 1 (different α's for different values of X)

• A version conditionalized on some background evidence e:  $\mathbf{P}(Y|X, \mathbf{e}) = \frac{\mathbf{P}(X|Y, \mathbf{e})\mathbf{P}(Y|\mathbf{e})}{\mathbf{P}(X|\mathbf{e})}$ 

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• A version conditionalized on some background evidence **e**:  $P(Y|X, e) = \frac{P(X|Y, e)P(Y|e)}{P(X|e)}$ 

 Used to assess diagnostic probability from causal probability:  $P(cause|effect) = \frac{P(effect|cause)P(cause)}{P(cause)}$ 

P(effect)

• *P*(*cause*|*effect*) goes from effect to cause (diagnostic direction) P(effect|cause) goes from cause to effect (causal direction)

$$\implies P(m|s) = \frac{P(s|m)P(m)}{P(s)} = \frac{0.7 \cdot 1/50000}{0.01} = 0.0014$$

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#### Example

 An expert doctor is likely to have causal knowledge ... P(symptoms|disease) (i.e., P(effect|cause))

• Ex: let *m* be meningitis, *s* be stiff neck

- P(m) = 1/50000, P(s) = 0.01 (prior knowledge, from statistics)
- "meningitis causes to the patient a stiff neck in 70% of cases":

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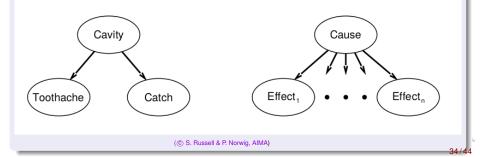
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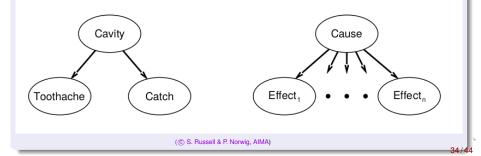
### Using Bayes' Rule: Combining Evidence

- A naive Bayes model is a probability model that assumes the effects are conditionally independent, given the cause
  - $\implies$  **P**(*Cause*, *Effect*<sub>1</sub>, ..., *Effect*<sub>n</sub>) = **P**(*Cause*)  $\prod_i$  **P**(*Effect*<sub>i</sub>|*Cause*)
    - total number of parameters is linear in n
    - ex: P(Cavity, Toothache, Catch) = P(Cavity)P(Toothache|Cavity)P(Catch|Cavity)
- Q: How can we compute P(Cause|Effect<sub>1</sub>,..., Effect<sub>k</sub>)?
   ex P(Cavity|toothache ∧ catch)?



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# Using Bayes' Rule: Combining Evidence [cont.]

Q: How can we compute  $P(Cause | Effect_1, ..., Effect_k)$ ?

• ex: P(Cavity|toothache ∧ catch)?

A: Apply Bayes' Rule

 $\mathbf{P}(Cavity | toothache \land catch)$ 

 $= \mathbf{P}(toothache \land catch|Cavity)\mathbf{P}(Cavity)/P(toothache \land catch)$ 

 $= lpha \mathbf{P}(toothache \land catch|Cavity) \mathbf{P}(Cavity)$ 

 $= lpha \mathbf{P}(toothache|Cavity)\mathbf{P}(catch|Cavity)\mathbf{P}(Cavity)$ 

•  $\alpha \stackrel{\text{def}}{=} 1/P(\text{toothache} \land \text{catch})$  not computed explicitly

General case:

 $\mathsf{P}(\mathit{Cause}|\mathit{Effect}_1,...,\mathit{Effect}_n) = lpha \mathsf{P}(\mathit{Cause}) \prod_i \mathsf{P}(\mathit{Effect}_i|\mathit{Cause})$ 

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 $\Rightarrow$  reduces from  $2^{n+1} - 1$  to 2n + 1 independent entries

## Using Bayes' Rule: Combining Evidence [cont.]

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• ex: P(Cavity|toothache ∧ catch)?

A: Apply Bayes' Rule

 $P(Cavity | toothache \land catch)$ 

- $= \mathbf{P}(\textit{toothache} \land \textit{catch}|\textit{Cavity})\mathbf{P}(\textit{Cavity})/P(\textit{toothache} \land \textit{catch})$
- $= \alpha \mathbf{P}(toothache \wedge catch|Cavity)\mathbf{P}(Cavity)$
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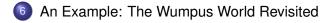
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## Outline



- 2 Basics on Probability
- Probabilistic Inference via Enumeration
- Independence and Conditional Independence
- 5 Applying Bayes' Rule



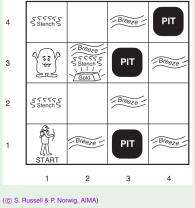
### A probability model of the Wumpus World

- Consider again the Wumpus World (restricted to pit detection)
- Evidence: no pit in (1,1), (1,2), (2,1), breezy in (1,2), (2,1)
- Q. Given the evidence, what is the probability of having a pit in (1,3), (2,2) or (3,1)?
  - Two groups of variables:
    - *P<sub>ij</sub> = true* iff [*i*, *j*] contains a pit ("causes")
    - *B<sub>ij</sub>* = *true* iff [*i*, *j*] is breezy ("effects", consider only *B*<sub>1,1</sub>, *B*<sub>1,2</sub>, *B*<sub>2,1</sub>)
  - Joint Distribution:

 $\mathbf{P}(P_{1,1},...,P_{4,4},B_{1,1},B_{1,2},B_{2,1})$ 

• Known facts (evidence):

•  $b^* \stackrel{\text{def}}{=} \neg b_{1,1} \land b_{1,2} \land b_{2,1}$ •  $p^* \stackrel{\text{def}}{=} \neg p_{1,1} \land \neg p_{1,2} \land \neg p_{2,1}$ 



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1,3	2,3	3,3	4,3
<sup>1,2</sup> B OK	2,2	3,2	4,2
1,1 OK	<sup>2,1</sup> B OK	3,1	4,1

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1,3	2,3	3,3	4,3
<sup>1,2</sup> B OK	2,2	3,2	4,2
1,1	<sup>2,1</sup> B	3,1	4,1
OK	OK		

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ОК	OK		

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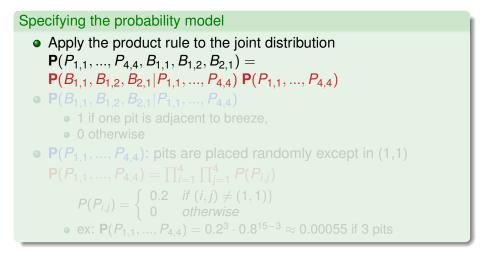
 $\mathbf{P}(P_{1,1},...,P_{4,4},B_{1,1},B_{1,2},B_{2,1})$ 

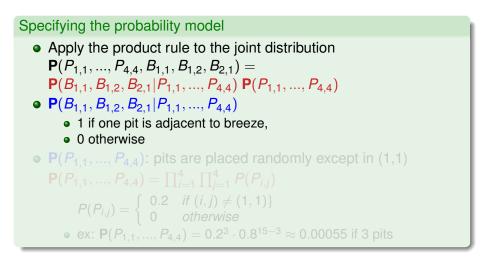
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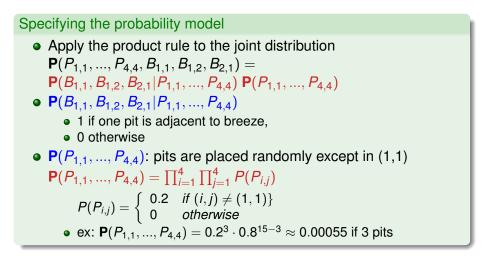
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1,1 OK	<sup>2,1</sup> B OK	3,1	4,1







#### Inference by enumeration

General form of query: P(Y|E = e) = αP(Y, E = e) = α Σ<sub>h</sub> P(Y, E = e, H = h) • Y: query vars; E,e: evidence vars/values; H,h: hidden vars/values
Our case: P(P<sub>1,3</sub>|p\*, b\*), s.t. the evidence is • b\* def ¬b<sub>1,1</sub> ∧ b<sub>1,2</sub> ∧ b<sub>2,1</sub> • p\* def ¬p<sub>1,1</sub> ∧ ¬p<sub>1,2</sub> ∧ ¬p<sub>2,1</sub>
Sum over hidden variables: P(P<sub>1,3</sub>|p\*, b\*) = • P(P<sub>1,3</sub>|p\*, b\*) = • P(P<sub>1,3</sub>|p\*, b\*) =

• *unknown* are all  $P_{ij}$ 's s.t. (*i*,*j*)  $\notin$  {(1, 1), (1, 2), (2, 1), (1, 3)}  $\Rightarrow 2^{16-4} = 4096$  terms of the sum!

Grows exponentially in the number of hidden variables H!
 Inefficient

#### Inference by enumeration

General form of query:  $P(Y|E = e) = \alpha P(Y, E = e) = \alpha \sum_{h} P(Y, E = e, H = h)$ • Y: query vars; E,e: evidence vars/values; H,h: hidden vars/values • Our case:  $\mathbf{P}(P_{1,3}|p^*, b^*)$ , s.t. the evidence is •  $b^* \stackrel{\text{def}}{=} \neg b_{1,1} \land b_{1,2} \land b_{2,1}$ •  $p^* \stackrel{\text{def}}{=} \neg p_{1,1} \land \neg p_{1,2} \land \neg p_{2,1}$ 2.3 3.3  $P_{13}$ <sup>1,2</sup> B OK 4,1 2,1 • unknown are all P<sub>ii</sub>'s s.t. в OK OK (© S. Russell & P. Norwig, AIMA) Grows exponentially in the number of hidden variables H!

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- Sum over hidden variables:
  - $\mathbf{P}(P_{1,3}|p^*,b^*) =$  $\alpha \sum_{unknown} \mathbf{P}(P_{1,3}|p^*, b^*, unknown)$
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<sup>(©</sup> S. Russell & P. Norwig, AIMA)

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#### Inference by enumeration

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$$b^* \stackrel{\text{def}}{=} \neg b_{1,1} \land b_{1,2} \land b_{2,1}$$

• 
$$p = p_{1,1} \wedge p_{1,2} \wedge p_{2,2}$$

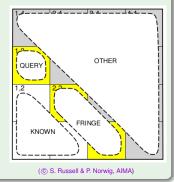
- Sum over hidden variables:
  - $P(P_{1,3}|p^*, b^*) = \alpha \sum_{unknown} P(P_{1,3}|p^*, b^*, unknown)$

• unknown are all  $P_{ij}$ 's s.t. (*i*, *j*)  $\notin$  {(1, 1), (1, 2), (2, 1), (1, 3)}  $\implies 2^{16-4} = 4096$  terms of the sum!

Grows exponentially in the number of hidden variables H!
 Inefficient

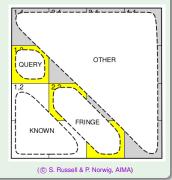
Using conditional independence

- Basic insight: Given the fringe squares (see below), *b*\* is conditionally independent of the other hidden squares
  - Unknown  $\stackrel{\text{def}}{=}$  Fringe  $\cup$  Other
- $\implies \mathbf{P}(b^*|p^*, P_{1,3}, Unknown) \stackrel{\text{def}}{=} \mathbf{P}(b^*|p^*, P_{1,3}, Fringe, Others) = \mathbf{P}(b^*|p^*, P_{1,3}, Fringe)$ 
  - Next: manipulate the query into a form where this equation can be used



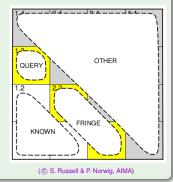
Using conditional independence

- Basic insight: Given the fringe squares (see below), *b*<sup>\*</sup> is conditionally independent of the other hidden squares
  - Unknown  $\stackrel{\text{def}}{=}$  Fringe  $\cup$  Other
- $\implies \mathbf{P}(b^*|p^*, P_{1,3}, \textit{Unknown}) \stackrel{\text{def}}{=} \mathbf{P}(b^*|p^*, P_{1,3}, \textit{Fringe}, \textit{Others}) = \mathbf{P}(b^*|p^*, P_{1,3}, \textit{Fringe})$ 
  - Next: manipulate the query into a form where this equation can be used



Using conditional independence

- Basic insight: Given the fringe squares (see below), *b*<sup>\*</sup> is conditionally independent of the other hidden squares
  - Unknown  $\stackrel{\text{def}}{=}$  Fringe  $\cup$  Other
- $\implies \mathbf{P}(b^*|p^*, P_{1,3}, Unknown) \stackrel{\text{def}}{=} \mathbf{P}(b^*|p^*, P_{1,3}, Fringe, Others) = \mathbf{P}(b^*|p^*, P_{1,3}, Fringe)$ 
  - Next: manipulate the query into a form where this equation can be used



 $\mathbf{P}(p^*, b^*) = P(p^*, b^*)$  is scalar; use as a normalization constant

 $\mathbf{P}(P_{1,3}|p^*,b^*) = \mathbf{P}(P_{1,3},p^*,b^*) / \mathbf{P}(p^*,b^*) = \underline{\alpha} \mathbf{P}(P_{1,3},p^*,b^*)$ 

#### Sum over the unknowns

$$\begin{split} \mathbf{P}(P_{1,3}|p^*,b^*) &= \mathbf{P}(P_{1,3},p^*,b^*) / \mathbf{P}(p^*,b^*) = \alpha \mathbf{P}(P_{1,3},p^*,b^*) \\ &= \alpha \sum_{\underline{unknown}} \mathbf{P}(P_{1,3},\underline{unknown},p^*,b^*) \end{split}$$

#### Use the product rule

$$\begin{aligned} \mathbf{P}(P_{1,3}|p^*,b^*) &= \mathbf{P}(P_{1,3},p^*,b^*) / \mathbf{P}(p^*,b^*) = \alpha \mathbf{P}(P_{1,3},p^*,b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(P_{1,3},unknown,p^*,\underline{b^*}) \\ &= \alpha \sum_{unknown} \mathbf{P}(\underline{b^*}|P_{1,3},p^*,unknown) \mathbf{P}(P_{1,3},p^*,unknown) \end{aligned}$$

### Separate unknown into fringe and other

$$\begin{split} \mathbf{P}(P_{1,3}|p^*,b^*) &= \mathbf{P}(P_{1,3},p^*,b^*) / \mathbf{P}(p^*,b^*) = \alpha \mathbf{P}(P_{1,3},p^*,b^*) \\ &= \alpha \sum_{\substack{unknown\\unknown}} \mathbf{P}(P_{1,3},unknown,p^*,b^*) \\ &= \alpha \sum_{\substack{unknown\\unknown}} \mathbf{P}(b^*|P_{1,3},p^*,\underline{unknown}) \mathbf{P}(P_{1,3},p^*,\underline{unknown}) \\ &= \alpha \sum_{\substack{fringe\\other}} \sum_{p} \mathbf{P}(b^*|p^*,P_{1,3},\underline{fringe},other) \mathbf{P}(P_{1,3},p^*,\underline{fringe},other) \end{split}$$

### b\* is conditionally independent of other given fringe

$$\begin{aligned} \mathbf{P}(P_{1,3}|p^*, b^*) &= \mathbf{P}(P_{1,3}, p^*, b^*) / \mathbf{P}(p^*, b^*) = \alpha \mathbf{P}(P_{1,3}, p^*, b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(P_{1,3}, unknown, p^*, b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(b^*|P_{1,3}, p^*, unknown) \mathbf{P}(P_{1,3}, p^*, unknown) \\ &= \alpha \sum_{fringe other} \mathbf{P}(b^*|p^*, P_{1,3}, \underline{fringe}, other) \mathbf{P}(P_{1,3}, p^*, fringe, other) \\ &= \alpha \sum_{fringe other} \mathbf{P}(b^*|p^*, P_{1,3}, \underline{fringe}) \mathbf{P}(P_{1,3}, p^*, fringe, other) \end{aligned}$$

#### Move $\mathbf{P}(b^*|p^*, P_{1,3}, fringe)$ outward

$$\begin{split} \mathbf{P}(P_{1,3}|p^*,b^*) &= \mathbf{P}(P_{1,3},p^*,b^*)/\mathbf{P}(p^*,b^*) = \alpha \mathbf{P}(P_{1,3},p^*,b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(P_{1,3},unknown,p^*,b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(b^*|P_{1,3},p^*,unknown)\mathbf{P}(P_{1,3},p^*,unknown) \\ &= \alpha \sum_{fringe other} \sum_{P(b^*|p^*,P_{1,3},fringe,other) \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe other} \sum_{p(b^*|p^*,P_{1,3},fringe)} \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \sum_{other} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3},p^*,fringe,other) \end{split}$$

#### All of the pit locations are independent

$$\begin{split} \mathbf{P}(P_{1,3}|p^*,b^*) &= \mathbf{P}(P_{1,3},p^*,b^*)/\mathbf{P}(p^*,b^*) = \alpha \mathbf{P}(P_{1,3},p^*,b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(P_{1,3},unknown,p^*,b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(b^*|P_{1,3},p^*,unknown) \mathbf{P}(P_{1,3},p^*,unknown) \\ &= \alpha \sum_{fringe other} \mathbf{P}(b^*|p^*,P_{1,3},fringe,other) \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe other} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3})P(p^*)P(fringe)P(other) \end{split}$$

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### Move $P(p^*)$ , $P(P_{1,3})$ , and P(fringe) outward

$$\begin{split} \mathbf{P}(P_{1,3}|p^*,b^*) &= \mathbf{P}(P_{1,3},p^*,b^*)/\mathbf{P}(p^*,b^*) = \alpha \mathbf{P}(P_{1,3},p^*,b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(P_{1,3},unknown,p^*,b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(b^*|P_{1,3},p^*,unknown) \mathbf{P}(P_{1,3},p^*,unknown) \\ &= \alpha \sum_{fringe} \sum_{other} \mathbf{P}(b^*|p^*,P_{1,3},fringe,other) \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \sum_{other} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3})P(p^*)P(fringe)P(other) \\ &= \alpha \underbrace{P(p^*)\mathbf{P}(P_{1,3})}_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \underbrace{P(fringe)}_{other} P(other) \end{split}$$

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### Remove $\sum_{other} P(other)$ because it equals 1

$$\begin{split} \mathbf{P}(P_{1,3}|p^*,b^*) &= \mathbf{P}(P_{1,3},p^*,b^*)/\mathbf{P}(p^*,b^*) = \alpha \mathbf{P}(P_{1,3},p^*,b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(P_{1,3},unknown,p^*,b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(b^*|P_{1,3},p^*,unknown)\mathbf{P}(P_{1,3},p^*,unknown) \\ &= \alpha \sum_{fringe} \sum_{other} \mathbf{P}(b^*|p^*,P_{1,3},fringe,other)\mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \sum_{other} \mathbf{P}(b^*|p^*,P_{1,3},fringe)\mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3})P(p^*)P(fringe)P(other) \\ &= \alpha P(p^*)\mathbf{P}(P_{1,3}) \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe)P(fringe) \\ &= \alpha P(p^*)\mathbf{P}(P_{1,3}) \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe)P(fringe) \end{split}$$

 $P(p^*)$  is scalar, so make it part of the normalization constant

$$\begin{split} \mathbf{P}(P_{1,3}|p^*,b^*) &= \mathbf{P}(P_{1,3},p^*,b^*)/\mathbf{P}(p^*,b^*) = \alpha \mathbf{P}(P_{1,3},p^*,b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(P_{1,3},unknown,p^*,b^*) \\ &= \alpha \sum_{unknown} \mathbf{P}(b^*|P_{1,3},p^*,unknown) \mathbf{P}(P_{1,3},p^*,unknown) \\ &= \alpha \sum_{\sum} \sum_{fringe \ other} \mathbf{P}(b^*|p^*,P_{1,3},fringe,other) \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe \ other} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3},p^*,fringe,other) \\ &= \alpha \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe) \sum_{other} \mathbf{P}(P_{1,3})P(p^*)P(fringe)P(other) \\ &= \alpha P(p^*)\mathbf{P}(P_{1,3}) \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe)P(fringe) \\ &= \frac{\alpha P(p^*)\mathbf{P}(P_{1,3}) \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe)P(fringe) \\ &= \frac{\alpha' \mathbf{P}(P_{1,3}) \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe)P(fringe) \\ &= \alpha' \mathbf{P}(P_{1,3}) \sum_{fringe} \mathbf{P}(b^*|p^*,P_{1,3},fringe)P(fringe) \end{split}$$

We have obtained:

**P**( $P_{1,3}|p^*, b^*$ ) =  $\alpha'$ **P**( $P_{1,3}$ )  $\sum_{fringe}$  **P**( $b^*|p^*, P_{1,3}, fringe$ ) *P*(fringe) ● We know that **P**( $P_{1,3}$ ) =  $\langle 0.2, 0.8 \rangle$ 

- We can compute the normalization coefficient  $\alpha'$  afterwards
- $\sum_{\text{fringe}} \mathbf{P}(b^* | p^*, P_{1,3}, \text{fringe}) P(\text{fringe})$ : only 4 possible fringes
- Start by rewriting as two separate equations:  $\mathbf{P}(p_{1,3}|p^*, b^*) = \alpha' P(p_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*, p_{1,3}, \text{fringe}) P(\text{fringe})$   $\mathbf{P}(\neg p_{1,3}|p^*, b^*) = \alpha' P(\neg p_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*, \neg p_{1,3}, \text{fringe}) P(\text{fringe})$

• We have obtained:

 $\mathbf{P}(P_{1,3}|p^*, b^*) = \alpha' \mathbf{P}(P_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*, P_{1,3}, \text{fringe}) P(\text{fringe})$ 

- We know that  $P(P_{1,3}) = (0.2, 0.8)$
- We can compute the normalization coefficient  $\alpha'$  afterwards
- $\sum_{fringe} \mathbf{P}(b^*|p^*, P_{1,3}, fringe) P(fringe)$ : only 4 possible fringes
- Start by rewriting as two separate equations:  $\mathbf{P}(p_{1,3}|p^*, b^*) = \alpha' P(p_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*, p_{1,3}, \text{fringe}) P(\text{fringe})$   $\mathbf{P}(\neg p_{1,3}|p^*, b^*) = \alpha' P(\neg p_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*, \neg p_{1,3}, \text{fringe}) P(\text{fringe})$

• We have obtained:

 $\mathbf{P}(P_{1,3}|p^*, b^*) = \alpha' \mathbf{P}(P_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*, P_{1,3}, \text{fringe}) P(\text{fringe})$ 

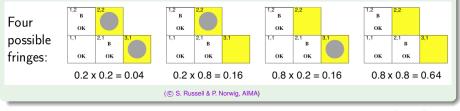
- We know that  $P(P_{1,3}) = (0.2, 0.8)$
- We can compute the normalization coefficient  $\alpha'$  afterwards
- $\sum_{\text{fringe}} \mathbf{P}(b^*|p^*, P_{1,3}, \text{fringe}) P(\text{fringe})$ : only 4 possible fringes
- Start by rewriting as two separate equations:  $P(p_{1,3}|p^*, b^*) = \alpha' P(p_{1,3}) \sum_{fringe} P(b^*|p^*, p_{1,3}, fringe) P(fringe)
   P(\neg p_{1,3}|p^*, b^*) = \alpha' P(\neg p_{1,3}) \sum_{fringe} P(b^*|p^*, \neg p_{1,3}, fringe) P(fringe)$

• We have obtained:

 $\mathbf{P}(P_{1,3}|p^*, b^*) = \alpha' \mathbf{P}(P_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*, P_{1,3}, \text{fringe}) P(\text{fringe})$ 

- We know that  $P(P_{1,3}) = \langle 0.2, 0.8 \rangle$
- We can compute the normalization coefficient  $\alpha'$  afterwards
- $\sum_{\text{fringe}} \mathbf{P}(b^* | p^*, P_{1,3}, \text{fringe}) P(\text{fringe})$ : only 4 possible fringes

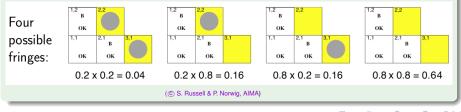
• Start by rewriting as two separate equations:  $\mathbf{P}(p_{1,3}|p^*, b^*) = \alpha' P(p_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*, p_{1,3}, \text{fringe}) P(\text{fringe})$   $\mathbf{P}(\neg p_{1,3}|p^*, b^*) = \alpha' P(\neg p_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*, \neg p_{1,3}, \text{fringe}) P(\text{fringe})$ 



• We have obtained:

 $\mathbf{P}(P_{1,3}|p^*, b^*) = \alpha' \mathbf{P}(P_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*, P_{1,3}, \text{fringe}) P(\text{fringe})$ 

- We know that  $P(P_{1,3}) = (0.2, 0.8)$
- We can compute the normalization coefficient  $\alpha'$  afterwards
- $\sum_{\text{fringe}} \mathbf{P}(b^*|p^*, P_{1,3}, \text{fringe}) P(\text{fringe})$ : only 4 possible fringes
- Start by rewriting as two separate equations:  $P(p_{1,3}|p^*, b^*) = \alpha' P(p_{1,3}) \sum_{fringe} P(b^*|p^*, p_{1,3}, fringe) P(fringe)$   $P(\neg p_{1,3}|p^*, b^*) = \alpha' P(\neg p_{1,3}) \sum_{fringe} P(b^*|p^*, \neg p_{1,3}, fringe) P(fringe)$



• Start by rewriting as two separate equations:  $\mathbf{P}(\begin{array}{c} p_{1,3}|p^*,b^*) = \alpha' P(\begin{array}{c} p_{1,3}) \sum_{\textit{fringe}} \mathbf{P}(b^*|p^*, \begin{array}{c} p_{1,3}, \textit{fringe}) P(\textit{fringe}) \\ \mathbf{P}(\neg p_{1,3}|p^*,b^*) = \alpha' P(\neg p_{1,3}) \sum_{\textit{fringe}} \mathbf{P}(b^*|p^*, \neg p_{1,3}, \textit{fringe}) P(\textit{fringe}) \end{array}$ 

• For each of them,  $P(b^*|...)$  is 1 if the breezes occur, 0 otherwise:  $\sum_{\text{fringe}} \mathbf{P}(b^*|p^*, p_{1,3}, \text{fringe})P(\text{fringe}) = 1.0.04 + 1.0.16 + 1.0.16 + 0 = 0.36$  $\sum_{\text{fringe}} \mathbf{P}(b^*|p^*, \neg p_{1,3}, \text{fringe})P(\text{fringe}) = 1.0.04 + 1.0.16 + 0 + 0 = 0.2$ 

 $\implies \mathbf{P}(P_{1,3}|p^*,b^*) = \alpha' \mathbf{P}(P_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*,P_{1,3},\text{fringe}) P(\text{fringe})$  $= \alpha' \langle 0.2, 0.8 \rangle \langle 0.36, 0.2 \rangle = \alpha' \langle 0.072, 0.16 \rangle$ 

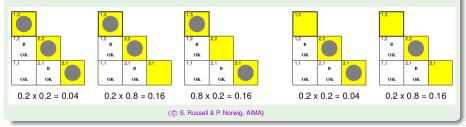
= (*normalization*, *s.t.*  $\alpha' \approx$  4.31)  $\approx$   $\langle$  0.31, 0.69 $\rangle$ 

• Start by rewriting as two separate equations:  $\mathbf{P}(\begin{array}{c} p_{1,3} | p^*, b^* ) = \alpha' P(\begin{array}{c} p_{1,3} ) \sum_{\textit{fringe}} \mathbf{P}(b^* | p^*, \begin{array}{c} p_{1,3}, \textit{fringe} ) P(\textit{fringe} ) \\ \mathbf{P}(\neg p_{1,3} | p^*, b^* ) = \alpha' P(\neg p_{1,3}) \sum_{\textit{fringe}} \mathbf{P}(b^* | p^*, \neg p_{1,3}, \textit{fringe}) P(\textit{fringe} ) \end{array}$ 

• For each of them,  $P(b^*|...)$  is 1 if the breezes occur, 0 otherwise:  $\sum_{\text{fringe}} P(b^*|p^*, p_{1,3}, \text{fringe})P(\text{fringe}) = 1.0.04 + 1.0.16 + 1.0.16 + 0 = 0.36$  $\sum_{\text{fringe}} P(b^*|p^*, \neg p_{1,3}, \text{fringe})P(\text{fringe}) = 1.0.04 + 1.0.16 + 0 + 0 = 0.2$ 

 $\implies \mathbf{P}(P_{1,3}|p^*,b^*) = \alpha' \mathbf{P}(P_{1,3}) \sum_{\text{fringe}} \mathbf{P}(b^*|p^*,P_{1,3},\text{fringe}) P(\text{fringe}) \\ = \alpha' \langle 0.2, 0.8 \rangle \langle 0.36, 0.2 \rangle = \alpha' \langle 0.072, 0.16 \rangle$ 

 $\alpha = (\textit{normalization}, \textit{ s.t. } lpha' pprox 4.31) pprox \langle 0.31, 0.69 
angle$ 



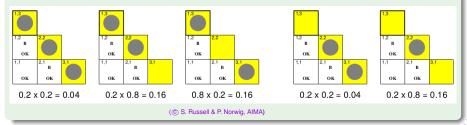
• Start by rewriting as two separate equations:

 $\begin{array}{l} \mathbf{P}( \ p_{1,3} | p^*, b^*) = \alpha' P( \ p_{1,3}) \sum_{\textit{fringe}} \mathbf{P}(b^* | p^*, \ p_{1,3}, \textit{fringe}) P(\textit{fringe}) \\ \mathbf{P}(\neg p_{1,3} | p^*, b^*) = \alpha' P(\neg p_{1,3}) \sum_{\textit{fringe}} \mathbf{P}(b^* | p^*, \neg p_{1,3}, \textit{fringe}) P(\textit{fringe}) \end{array}$ 

• For each of them,  $P(b^*|...)$  is 1 if the breezes occur, 0 otherwise:  $\sum_{fringe} \mathbf{P}(b^*|p^*, p_{1,3}, fringe)P(fringe) = 1.0.04 + 1.0.16 + 1.0.16 + 0 = 0.36$  $\sum_{fringe} \mathbf{P}(b^*|p^*, \neg p_{1,3}, fringe)P(fringe) = 1.0.04 + 1.0.16 + 0 + 0 = 0.2$ 

 $\implies \mathbf{P}(P_{1,3}|\boldsymbol{p}^*, \boldsymbol{b}^*) = \alpha' \mathbf{P}(P_{1,3}) \sum_{\textit{fringe}} \mathbf{P}(\boldsymbol{b}^*|\boldsymbol{p}^*, P_{1,3}, \textit{fringe}) P(\textit{fringe}) \\ = \alpha' \langle 0.2, 0.8 \rangle \langle 0.36, 0.2 \rangle = \alpha' \langle 0.072, 0.16 \rangle$ 

= (normalization, s.t.  $\alpha' \approx$  4.31)  $\approx \langle 0.31, 0.69 \rangle$ 



Compute  $\mathbf{P}(P_{2,2}|p^*, b^*)$  in the same way.