

# Basic Notions of Maximum Likelihood Estimation and Regression

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The basic idea of MLE is simple

- Given I observed event  $B$ , what is the probability the event  $A$  occurred?
- Also: Given I have the sample  $\{X_i\}$  what is the most likely population / process that generated it?
- MLE under certain hypotheses can be shown to be asymptotically optimum
- For small sample sets the estimation can be biased and give wrong results
- Unless there are some additional strong constraints MLE can be computationally very heavy
  - There are no “general” closed form solutions
  - If the state space of  $A$  is continuous, then we can in general only have an approximate solution

MLE is based on Bayes' Theorem

$$P[B_j|A] = \frac{P[A|B_j]P[B_j]}{P[A]} \Leftrightarrow P[A] = \frac{P[A|B_j]P[B_j]}{P[B_j|A]}$$

- MLE maximizes the a-posteriori probability of a conditional probability
- The maximization is done on some parameters of the conditioning events

Let  $\{X_i; i = 1, 2, \dots, n\}$  be a sample set and  $\Theta = \{\theta_1, \theta_2, \dots, \theta_k\}$  be a set or vector of parameters to be estimated  
Define a likelihood function  $L(\Theta)$  as:

$$L(\Theta) = \mathbf{P}[X_1 = x_1, X_2 = x_2, \dots, X_n = x_n | \Theta]$$

if the population is described by a discrete PMF

or

$$L(\Theta) = f_X(x | \Theta)$$

if the population is described by a continuous pdf



Now the problem is trivial: find  $\Theta$  such that  $L(\Theta)$  is maximum

In math

$$\hat{\Theta} : \operatorname{argmax}_{\Theta} L(\Theta)$$



- We need to know the joint probability of  $n$  random variables
- If they are not i.i.d. ... **game over!**
- If we know the sample set is i.i.d. then the likelihood functions reduce to

$$L(\Theta) = \prod_{i=1}^n \mathbf{P}[X_i = x_i | \Theta]$$

if the population is described by a discrete PMF, or

$$L(\Theta) = \prod_{i=1}^n f_{X_i}(x_i | \Theta)$$

if the population is described by a continuous pdf



In case of i.i.d. sets (& some other cases), as the likelihood function  $L(\Theta)$  is described as a product it is custom to use logarithmic likelihood function  $l(\Theta) = \log[L(\Theta)]$  so that the maximization problem is described by a sum and not by a product



$$l(\Theta) = \sum_{i=1}^n \mathbf{P}[X_i = x_i | \Theta]$$

if the population is described by a discrete PMF, or

$$l(\Theta) = \sum_{i=1}^n f_{X_i}(x_i | \Theta)$$

if the population is described by a continuous pdf



- Depending on  $\Theta$  the problem can still be computationally very difficult (even in i.i.d. cases)
- Under some fairly general conditions of regularity of both the distributions and the  $\Theta$  parameter set, then the optimization, in general an NP-complete problem, can be reduced to a set of  $k$  joint partial differential equations, where finding the zeros may be easy (?!?)

$$\frac{\delta L(\Theta)}{\delta \theta_i}; \quad i = 1, 2, \dots, k$$

- Really the only case where MLE is simple and works without hassles is when  $\theta_i$  are orthogonal and the partial differential equations either reduce to normal differential equations or we can in any case apply the gradient algorithm





- Really the only case where MLE is simple and works without hassles is when  $\theta_i$  (the set of parameters) are orthogonal and the partial differential equations either reduce to normal differential equations

$$\frac{dL(\Theta)}{d\theta_i}; \quad i = 1, 2, \dots, k$$

- or we can in any case apply the gradient algorithm (only one minimum exists)



- For instance if  $\{X_i\}$  is drawn from a gamma distribution and  $\theta_1$  and  $\theta_2$  are the parameters  $\lambda$  and  $\alpha$  of the distribution, then the set of 2 partial differential equations have no closed form solution and we have to resort to numerical methods (that's why you find the function in Matlab!!)

- For another totally “casual” example, if  $\{X_i\}$  is drawn from a gamma distribution affected by random Gaussian noise samples  $Y_i$  distributed as  $N(0, \sigma)$  and  $\theta_1$ ,  $\theta_2$  and  $\theta_3$  are the parameters  $\lambda$ ,  $\alpha$ , and  $\sigma$  of the two distributions, then we have to compute the distribution of

$$Z_i = X_i + Y_i$$

$$f_Z(z) = f_X(x) * f_Y(y)$$

where  $*$  is the convolutional product so

$$f_Z(z) = \int_{-\infty}^{\infty} \frac{\lambda^\alpha t^{\alpha-1} e^{-\lambda x}}{\Gamma(\alpha)} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-z)^2}{2\sigma^2}} dx$$

and there is no solution to the MLE, unless we resort to (complex) numerical methods



MLE is instead simple when  $\Theta$  is a partition of a probability space or a finite set of deterministic conditions. For example, it is the base for optimal detection in digital communications

- The key “problem” of digital transmission is finding the best strategy to decide what symbol  $S_i(t)$  has been transmitted given we have received a symbol  $R(t)$
- Find the maximum over  $j$  of

$$P[S_j|R] = \frac{P[R|S_j]P[S_j]}{P[R]}$$

- $R(t)$  can be modeled as  $R(t) = S_i(t) + N(0, \sigma)$

$$P[S_j | (S_j(t) + N(0, \sigma))] = \frac{P[(S_j(t) + N(0, \sigma)) | S_j] P[S_j]}{P[(S_j(t) + N(0, \sigma))]}$$

- Thus the MLE problem is reduced to a minimum distance problem

$$\min_j (||S_j - R||)$$

- More reasoning at the blackboard.



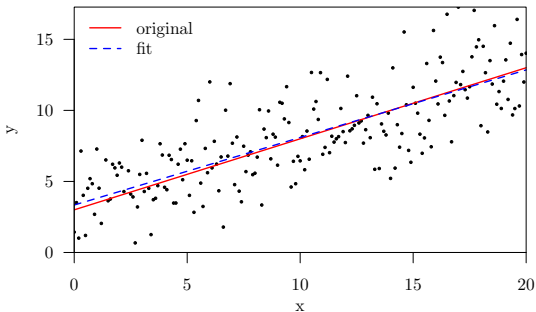
- Consider two joint RV  $X, Y$  and a dependence function  $d(\cdot)$  such that  $Y = d(X) + \epsilon$  where  $\epsilon$  is a residual error
- Our problem is finding  $d(\cdot)$  such that  $d(X)$  is as close as possible to  $Y$  in some appropriate sense, e.g., minimizing a euclidean distance or a generic norm such as  $l_\infty$  or any proper measure
- Let  $D = Y - d(X)$  be the random variable that measures the residual error done because we do not know  $f_{X,Y}(x, y)$ , and we approximate the dependence with the function  $d(\cdot)$
- The most common measure of the difference is  $E[D^2]$

- The function  $d(x)$  that minimizes  $E[D^2]$  is called the **Least-square regression curve**
- It is not difficult to show that this function is  $d(x) = E[Y|x]$
- However the conditional distribution  $f_{Y|x}(y|x)$  is normally very difficult to find
- It is common practice to limit the structure of  $d(x)$  (e.g., to a polynomial function) to make the problem more tractable





A scatter diagram is nothing else than an  $(x, y)$  plot of the outcome of  $n$  random experiments on the pair  $X, Y$



Scatter diagram with the linear regression of the points and the “true” linear relationship



- The simplest form of dependence is assuming that the function is linear:  $d(x) = a + bx$
- Clearly this is a huge limitation to the dependence relationship, but in many cases it is useful and it can be treated easily
- In this case the problem of finding the optimal fitting curve reduces to minimize the following

$$G(a, b) = e[D^2] = E[(Y - d(X))^2] = E[(Y - a - bX)^2]$$

- Let  $\mu_x, \mu_y, \sigma_x^2, \sigma_y^2$  be the mean and variance of  $X$  and  $Y$  respectively, and also  $\rho = \frac{\text{Cov}(X, Y)}{\sigma_x \sigma_y}$

- Then expanding  $G(a, b)$  yields

$$\begin{aligned} G(a, b) &= \sigma_y^2 + b^2 \sigma_x^2 + (\mu_y - a)^2 + b^2 \mu_x^2 - 2b\rho\sigma_x\sigma_y \\ &\quad - 2b\mu_x(\mu_y - a) \\ &= \sigma_y^2 + b^2 \sigma_x^2 + (\mu_y - a - b\mu_x)^2 - 2b\rho\sigma_x\sigma_y \end{aligned}$$

- To find the minimum of  $G(a, b)$  we have to find the point where the partial derivatives with respect to  $a$  and  $b$  are zero



$$\frac{\delta G(a, b)}{\delta a} = -2(\mu_y - a - b\mu_x) = 0$$

$$\frac{\delta G(a, b)}{\delta b} = 2b\sigma_x^2 - 2\mu_x(\mu_y - a - b\mu_x) - 2\rho\sigma_x\sigma_y = 0$$

Solving the equations we find that the optimal values of  $a$  and  $b$  are

$$b = \rho \frac{\sigma_y}{\sigma_x}$$
$$a = \mu_y - b\mu_x$$

You normally find subroutines and function to perform a linear regression in any statistical tool



- If the relationship is not linear, then finding the regression can be very difficult, even if the polynomial structure is given (it is not like the deterministic case of fitting)
- The exception is the exponential relation

$$Y = ae^{bX}$$

where we can simply take the logarithm and do a linear fitting of the logarithm